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NOTE ON A GENERALIZATION  
OF A LAW OF LARGE NUMBERS

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## NOTE ON A GENERALIZATION OF A LAW OF LARGE NUMBERS

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### ABSTRACT

A sequence of probability distribution functions controlling behavior of a sequence of mutually independent random variables is given as an outcome of a "generator" stochastic process. Even if the law of large numbers holds for each outcome of the generator process, the limit may depend on the outcome. In this paper a sufficient condition where the limits in the law of large numbers are constant for almost all outcomes of the generator process is given.

### § 1. Introduction

Let us consider Bernoulli trials, i.e., repeated independent trials such that there are only two possible outcomes for each trial, say, "success" and "failure", and their probabilities remain the same throughout the trials. It is usual to denote the two probabilities by  $p$  and  $q(=1-p)$ . If  $S_n$  is the number of successes in  $n$  trials, then  $S_n/n$  is the frequency of successes and asymptotically tends to  $p$  as  $n$  increases. Accurately, *as  $n$  increases, the probability that the frequency of successes deviates from  $p$  by more than any preassigned  $\epsilon > 0$  tends to zero, i.e.*

$$P\left\{\left|\frac{S_n}{n} - p\right| < \epsilon\right\} \rightarrow 1, \text{ as } n \rightarrow \infty.$$

This is the weak law of large numbers due to J. BERNOULLI (1713). E. BOREL (1909) made a statement much stronger than the Bernoulli's, which has been known as the strong law of large numbers: *For almost every sequence of trials the frequency of successes tends to  $p$ , as  $n$  increases, i.e.*

$$P\left\{\lim_{n \rightarrow \infty} \frac{S_n}{n} = p\right\} = 1.$$

A. N. KOLMOGOROV (1928, 1929) obtained a generalization of these two theorems. *If the variances of the random variables  $X_k$  are bounded by one and the same constant, or more generally if  $\sum_{n=1}^{\infty} \text{Var}(X_n)/n^2 < \infty$ , then the sequence of mutually independent random variables  $X_1, X_2, \dots$  obeys the strong law of large numbers, i.e.*

$$\frac{1}{n} \sum_{k=1}^n X_k - \frac{1}{n} \sum_{k=1}^n EX_k \longrightarrow 0, \text{ as } n \longrightarrow \infty$$

holds with probability one.

In this paper we consider one generalization of Kolmogorov's theorem stated above. Suppose that a sequence of probability distribution functions, which governs behavior of a sequence of mutually independent random variables (abbreviated a seq. of mirv.'s) is not given *a priori* but is determined by an outcome of a stochastic process called, a generator process. Even if the law of large numbers holds for each outcome of the generator process, the limit generally depends on the outcome. In this paper a sufficient condition imposed on the generator process is derived so that the limits are the same for almost all outcomes of the generator process.

## § 2. The main theorem.

Let us introduce some notations and definitions.

Let  $I = \{1, 2, 3, \dots\}$ ,  $I_0 = \{0, 1, 2, \dots\}$ ,  $R = (-\infty, \infty)$  and  $S = \{F_i, i \in I\}$ , where  $F_i$ 's are probability distribution functions. Put  $W_1 = (I)^{I_0}$  and  $W_2 = (R)^{W_1}$ . Symbols  $BW_1$  and  $BW_2$  denote the smallest  $\sigma$ -algebras containing cylinder subsets of  $W_1$  and Borel cylinder subsets of  $W_2$  respectively. Here we define two shift-transformations  $T_1: W_1 \rightarrow W_1$  and  $T_2: W_2 \rightarrow W_2$  by

$$W_1 \ni w_1 = (i_0, i_1, i_2, \dots) \longrightarrow T_1 w_1 = (i_1, i_2, i_3, \dots) \in W_1,$$

and

$$\begin{aligned} W_2 \ni w_2 = ((x_1^{(0)}, x_2^{(0)}, \dots), (x_1^{(1)}, x_2^{(1)}, \dots), \dots) \longrightarrow \\ T_2 w_2 = ((x_1^{(1)}, x_2^{(1)}, \dots), (x_1^{(2)}, x_2^{(2)}, \dots), \dots) \in W_2 \end{aligned}$$

respectively.

Now let us introduce a probability measure  $P_1$  on a measurable space  $(W_1, BW_1)$  and define  $X_n(\cdot)$  on  $W_1$  as a coordinate function, i.e.  $X_n(w_1) = i_n$  for  $w_1 = (i_0, i_1, \dots, i_n, \dots)$ . Let us call the sequence of random variables  $\{X_n(\cdot); n \geq 0\}$  a generator process. Note that if  $T_1$  is measure-preserving, then for every  $i \in I$ ,

$$P_1\{w_1: X_n(w_1) = i\} = P_1\{w_1: X_0(w_1) = i\}, \text{ for } n \geq 0.$$

Let  $P_2$  be a probability measure on  $(W_2, BW_2)$  satisfying the condition that for every integers  $k, n_1 < n_2 < \dots < n_k, i_1 < i_2 < \dots < i_k$  and reals  $x_1, x_2, \dots, x_k$ ,

$$\begin{aligned} P_2\{w_2: Y_{n_1}(i_1, w_2) \leq x_1, Y_{n_2}(i_2, w_2) \leq x_2, \dots, Y_{n_k}(i_k, w_2) \leq x_k\} \\ = F_{i_1}(x_1) F_{i_2}(x_2) \dots F_{i_k}(x_k), \end{aligned}$$

where  $F_{i_j} \in \mathcal{S}$  and  $Y_{n_j}(i_j, \cdot)$  is a coordinate function on  $W_2$  such that

$$Y_{n_j}(i_j, w_2) = x_{i_j}^{(n_j)}$$

for  $w_2 = ((x_1^{(0)}, x_2^{(0)}, \dots), \dots, (x_1^{(n_j)}, x_2^{(n_j)}, \dots), \dots) \in W_2$ . Thus  $\{Y_n(i, \cdot): n \geq 0, i \geq 1\}$  is a seq. of mirv.'s on  $(W_2, BW_2, P_2)$ . Note that the shift-transformation  $T_2$  is measure-preserving and mixing. Let  $(W, BW, P)$  be a product probability space of  $(W_1, BW_1, P_1)$  and  $(W_2, BW_2, P_2)$ . We define a seq. of r.v.'s  $\{Z_n(\cdot): n \geq 0\}$  on  $(W, BW, P)$  and a transformation  $T: W \rightarrow W$  such that

$$Z_n(w) = Z_n(w_1, w_2) = Y_n(X_n(w_1), w_2)$$

and

$$Tw = T(w_1, w_2) = (T_1 w_1, T_2 w_2)$$

respectively.

Under the assumptions stated above, we have the following theorem.

**THEOREM.** *If  $T_1$  is measure-preserving and ergodic, and if  $\sum_{i \in I} |m_i| p_i < \infty$ , then for a.a.  $w$  (almost all  $w$ ),*

$$\frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) \longrightarrow EZ_0 = \int_W Z_0(w) P(dw), \text{ as } n \longrightarrow \infty,$$

where

$$m_i = \int_{-\infty}^{\infty} x dF_i(x) \text{ and } p_i = P_1\{w_1: X_0(w_1) = i\} \text{ for } i \in I.$$

*Proof.* From measure-preservation of  $T_1$  and  $T_2$  it follows that for every  $n \geq 0$ ,  $E|Z_n| = \sum_{i \in I} E|Y_n(i, w_2)| P_1\{w_1: X_n(w_1) = i\} = \sum_{i \in I} |m_i| p_i < \infty$ . Thus by Birkhoff's individual ergodic theorem (c.f. [3]) there exists a  $T$ -invariant and integrable function  $\hat{Z}(w)$  such that, for a.a.  $w$

$$\frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) \longrightarrow \hat{Z}(w), \text{ as } n \longrightarrow \infty,$$

and  $EZ_k = E\hat{Z}$  for  $k \geq 0$ . Since  $T_1$  is ergodic and  $T_2$  is mixing,  $T$  is ergodic. This can be proved in the same way as the mixing theorem (c.f. [3]).

It follows from this that  $\hat{Z}(w) = \text{const.}$  for a.a.  $w$ . Therefore  $EZ_0 = E\hat{Z} = \text{const.} = \hat{Z}(w)$  for a.a.  $w$  and this completes the proof.

**COROLLARY.** *If the same assumptions as in the theorem are satisfied and if also  $E|Z_0|^2 < \infty$ , then*

$$E \left| \frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) - EZ_0 \right|^2 \longrightarrow 0, \text{ as } n \longrightarrow \infty.$$

*Proof.* Since  $T$  is measure-preserving, we have  $E|Z_n|^2 = E|Z_0|^2 < \infty$  for  $n \geq 0$ . It follows from this by the mean ergodic theorem due to J. von NEUMANN (1932) that there exists a  $T$ -invariant random variable  $Z^*(w)$  such that

$$E \left| \frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) - Z^*(w) \right|^2 \longrightarrow 0, \text{ as } n \longrightarrow \infty.$$

This implies that as  $n \rightarrow \infty$ ,

$$\frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) \longrightarrow Z^*(w)$$

in probability. Choosing a subsequence  $\{n^*\} \subset \{n\}$ , we have that for a.a.w

$$\frac{1}{n^*} \sum_{k=0}^{n^*-1} Z_k(w) \longrightarrow Z^*(w), \text{ as } n^* \longrightarrow \infty.$$

Hence by the theorem stated above, we must have  $Z^*(w) = \hat{Z}(w)$  a.e. This completes the proof.

### § 3. Example

Consider a Markov chain as a generator process. Let  $I = \{0, 1, 2, \dots, N\}$ ,  $I_0 = \{0, 1, 2, \dots\}$ ,  $R = \{0, 1\}$  and  $S = \left\{ F_i = \left( \frac{i}{N}, 1 - \frac{i}{N} \right), i \in I \right\}$ . We interpret  $(W_1, BW_1)$  and  $(W_2, BW_2)$  as the same sense in § 2. Let also

$$M = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & \cdot & \cdot & \cdot & N-2 & N-1 & N \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ \cdot \\ \cdot \\ \cdot \\ N-2 \\ N-1 \\ N \end{matrix} & \left( \begin{matrix} 1-c & c & 0 & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & 0 \\ 1-c & 0 & c & 0 & \cdot & \cdot & \cdot & \cdot & \cdot & 0 \\ 0 & 1-c & 0 & c & 0 & \cdot & \cdot & \cdot & \cdot & 0 \\ 0 & 0 & 1-c & 0 & c & 0 & \cdot & \cdot & \cdot & 0 \\ \cdot & \cdot & \cdot & 0 & 1-c & 0 & c & 0 & \cdot & 0 \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & 0 \\ \cdot & \cdot & \cdot & \cdot & \cdot & 0 & 1-c & 0 & c & 0 \\ 0 & \cdot & \cdot & \cdot & \cdot & \cdot & 0 & 1-c & 0 & c \\ 0 & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & 0 & 1-c & c \end{matrix} \right), \end{matrix} \quad (14)$$

where  $0 < c < 1$ .

A normalized non-negative solution of the system of equations

$$vM = v \quad (15)$$

is  $p = (p_i: 0 \leq i \leq N)$ , where

$$\begin{aligned} p_i &= \frac{1+r}{1-r^{N+1}} r^i, \text{ when } r \neq 1 \\ &= \frac{1}{N+1}, \text{ when } r = 1, \end{aligned} \quad (16)$$

and

$$r = \frac{c}{1-c}.$$

Let  $\{X_n(\cdot); n \geq 0\}$  be a Markov chain, which is defined by the initial distribution  $p$  and the transition matrix  $M$ , on  $(W, BW, P_1)$  with the state space  $S$ . It follows from  $pM=p$  that a shift-transformation  $T_1: W_1 \rightarrow W_1$  defined by

$$X_n(T_1 w_1) = X_{n+1}(w_1) \quad (n \geq 0)$$

is measure-preserving and ergodic. Let  $P_2$  be a probability measure on  $(W_2, BW_2)$  satisfying the condition that for every integers  $k, n_1 < n_2 < \dots < n_k, i_1 < i_2 < \dots < i_k$  and  $e_1, e_2, \dots, e_k$

$$P_2\{w_2: Y_{n_j}(i_j, w_2) = e_j, 1 \leq j \leq k\} = \prod_{j=1}^k X(i_j, e_j),$$

where  $X(i_j, e_j) = i_j/N$ , if  $e_j = 1$  and  $1 - i_j/N$ , if  $e_j = 0$ , and  $Y_{n_j}(i_j, \cdot)$  is a coordinate function on  $W_2$ . Transformations  $T_2$  and  $T$ , a probability space  $(W, BW, P)$  and a seq. of r.v.'s  $\{Z_n(\cdot); n \geq 0\}$  are understood to be defined in the same way as in § 2. Then by the theorem in § 2, as  $n \rightarrow \infty$ , for a.a.w

$$\begin{aligned} \frac{1}{n} \sum_{k=0}^{n-1} Z_k(w) &\longrightarrow \sum_{i \in I} p_i \frac{i}{N} = \left( \frac{r(1-r^N)}{N(1-r)^2} - \frac{r^{N+1}}{1-r} \right) \frac{1+r}{1-r^{N-1}}, \text{ when } r \neq 1 \\ &= \frac{1}{2}, \text{ when } r = 1. \end{aligned}$$

This relation also holds in the mean convergence sense as justified by the corollary to the theorem stated above.

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